

# **An essay on mixed-frequency data, aggregated time series, and causality.**

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### **Resumo**

Entender as complexidades das relações econômicas é crucial para formuladores de políticas, pesquisadores e analistas. A agregação temporal, onde a frequência de geração de dados excede a frequência de coleta de dados, apresenta desafios significativos na análise econômica. Essa discrepância pode levar a realizações não observ´aveis do processo estoc´astico original, afetando as propriedades dos dados de séries temporais. Abordar esses desafios é vital para detectar e interpretar com precisão as relações causais entre variáveis econômicas. Nossa pesquisa visa identificar como a agregação temporal pode interferir na detecção de causalidade entre séries temporais. Tamb´em demonstramos como um teste de causalidade Sims modificado pode ser empregado para detectar causalidade em modelos de frequências mistas. Nossas simulações de Monte Carlo mostram boas propriedades de tamanho e poder para amostras finitas. Finalmente, testamos a causalidade entre o PIB dos EUA e indicadores macroeconômicos mensais.

Palavras-Chave: Séries temporais, Agregação temporal, Causalidade de Granger, Causalidade de Sims.

# **Abstract**

Temporal aggregation, where the data generation frequency exceeds the data collection frequency, poses significant challenges in economic analysis. This discrepancy can lead to unobservable realizations of the original stochastic process, which in turn affects the properties of time series data. Consequently, addressing these challenges is crucial for accurately detecting and interpreting causal relationships between economic variables. In our research, we aim to identify how temporal aggregation can interfere with the detection of causality between time series. Furthermore, we demonstrate the application of a modified Sims causality test to detect causality in mixed-frequency models. Our Monte Carlo simulations indicate that this test exhibits good finite sample size and power properties. Finally, we apply our methodology to test the causality between U.S. GDP and macroeconomic monthly indicators.

**Keywords:** Time series, Temporal aggregation, Granger causality, Sims causality.

# **Contents**



### <span id="page-6-0"></span>**1 Introduction**

Understanding the intricacies of economic relationships is crucial for policymakers, researchers, and analysts. Temporal aggregation, where the data generation frequency exceeds the data collection frequency, poses significant challenges in economic analysis, as described in [Zellner e Montmarquette](#page-30-0) [\(1971\)](#page-30-0). This discrepancy can lead to unobservable realizations of the original stochastic process, affecting the properties of time series data. Consequently, addressing these challenges is vital to accurately detecting and interpreting causal relationships between economic variables.

In this work, our first goal is to understand how temporal aggregation affects the detection of causality between two time series. Secondly, we aim to propose a method for correctly testing causality in situations where spurious causality exists. To achieve this, we demonstrate the existence or absence of spurious causality using four theoretical models considering a bivariate time series model. We then show that, with a slight modification, the Sims causality methodology is suitable for detecting causality when one series is aggregated and the other is not.

As stated in [Marcellino](#page-29-0) [\(1999\)](#page-29-0), temporal aggregation occurs when the frequency of data generation is higher than the frequency of data collection, resulting in some realizations of the original stochastic process being unobservable. The author also provides a detailed examination of which properties of the disaggregated time series remain invariant after aggregation and which do not, referring to this discrepancy as temporal aggregation bias. Further instances of common estimation problems associated with time aggregations can be found in [Weiss](#page-30-1) [\(1984\)](#page-30-1) and [Swanson e](#page-29-1) [Granger](#page-29-1) [\(1997\)](#page-29-1).

Granger first introduced the causality concepts that became known as Granger causality in [Granger](#page-28-0) [\(1963\)](#page-28-0). Although Granger's concept relates to the ability of one time series to predict another, conditional on a given information set, [Chalak e](#page-28-1) [White](#page-28-1) [\(2012\)](#page-28-1) demonstrates that Granger's concept is closely linked with the causal notions of the Pearl Causal Model discussed in [Pearl](#page-29-2) [\(2009\)](#page-29-2). Later, several empirical investigations, such as those in [Weiss](#page-30-1) [\(1984\)](#page-30-1), demonstrated that causality properties are not invariant to temporal aggregation. This can lead to spurious conclusions about the relationships between time series, a phenomenon referred to as spurious causality. To address this issue, [McCrorie e Chambers](#page-29-3) [\(2006\)](#page-29-3) proposed formulating models in continuous time to correct the effects of temporal aggregation in observed discrete data through a discrete-time analog. Similarly, [Renault et al.](#page-29-4) [\(1998\)](#page-29-4) utilized a continuous-time model to distinguish between true and spurious causality. Moreover, [Breitung e Swanson](#page-28-2) [\(2002\)](#page-28-2) observed that causality relationships seem to change when moving to a finer sampling interval. For more examples of the effects of temporal aggregation on causality testing, see [Rajaguru e Abeysinghe](#page-29-5) [\(2010\)](#page-29-5) and [Xu](#page-30-2) [\(1996\)](#page-30-2).

An alternative approach to identifying causality between non-aggregated observations was introduced in [Sims](#page-29-6) [\(1972\)](#page-29-6). Sims suggested that in a regression analysis if causality only moves from current and past values of exogenous variables to an endogenous variable, the coefficients for future values of the exogenous variables should be zero. This approach, known as the Sims causality test, has been widely applied. For example, in [Macunovich e Easterlin](#page-29-7) [\(1988\)](#page-29-7), Granger-Sims causality tests were applied to monthly age-specific data, demonstrating the technique's value in pinpointing the effective lag between business cycles and fertility in the United States. Similarly, [Chow](#page-28-3) [\(1987\)](#page-28-3) explored the causal link between export growth and industrial development in eight Newly Industrializing Countries using Sims' causality test to reveal a robust bidirectional causality between export growth and industrial development in most of them. For further applications of the Sims method, see [Heckman](#page-28-4) [\(2000\)](#page-28-4) and [Holland](#page-29-8) [\(1986\)](#page-29-8).

The paper is organized as follows: Section [2](#page-8-0) explores the characteristics of Granger causality; Section [3](#page-11-0) illustrates how aggregation leads to spurious causality; Section [4](#page-17-0) introduces the proposed modified Sims causality test; Section [5](#page-19-0) discuss the technical implementation details; Section [6](#page-20-0) reports the finite sample behavior of the proposed Sims test; Section [7](#page-25-0) presents an empirical application with quarterly GDP and monthly US indicators; Section [8](#page-27-0) concludes the paper and the Appendix showcases the intermediate results and expands upon the results presented in Section [3](#page-11-0) using the same aggregation pattern as presented in the GDP series.

### <span id="page-8-0"></span>**2 Granger Causality**

This section focuses on Granger's proposed measurement of causality between two variables, introduced in [Granger](#page-28-5) [\(1969\)](#page-28-5). [Hamilton](#page-28-6) [\(1994\)](#page-28-6) definition based on Vector Autoregression models (VARs) highlights the forecasting effectiveness of certain variables on others. Under this approach, the researcher aims to determine if lagged observations of the series  $\{x_t\}_{t=0}^{\infty}$  can contribute to forecasting the series  $\{y_t\}_{t=0}^{\infty}$ . If not, it can be concluded that the past values of  $x_t$  do not Granger-cause present values of  $y_t$ . This concept is denoted as  $x_t$  not Granger-causing  $y_t$  if:

<span id="page-8-1"></span>
$$
\text{MSE}[E(y_t|y_{t-1}, y_{t-2}, \ldots)] = \text{MSE}[E((y_t|y_{t-1}, y_{t-2}, \ldots, x_{t-1}, x_{t-2}, \ldots)],
$$

which means, for the linear case, that the past of  $x_t$  lacks explanatory power for  $y_t$ whenever the mean squared error (MSE) of forecasting  $y_t$ , conditioned exclusively on its historical values, is statistically equivalent to the MSE of forecasting *y<sup>t</sup>* using both its own and  $x_t$ 's history.

According to [Hamilton](#page-28-6) [\(1994\)](#page-28-6), considering past values of  $y_t$  and  $x_t$  and assuming a lag length *p*, we can define the augmented autoregressive model as:

$$
y_t = c + \alpha_1 y_{t-1} + \dots + \alpha_p y_{t-p} + \beta_1 x_{t-1} + \dots + \beta_p x_{t-p} + v_t \tag{2.1}
$$

where the error term  $v_t$  is assumed to be independent and identically distributed. In other words, as stated in [Kuersteiner](#page-29-9) [\(2010a\)](#page-29-9), testing for Granger causality involves assessing whether the coefficients associated with lags of  $x_t$  in the  $y_t$  series are statistically equal to zero. If this condition is not rejected, it implies that  $x_t \nrightarrow y_t$ .

As shown in [Marcellino](#page-29-0) [\(1999\)](#page-29-0), testing for Granger causality between aggregated time series variables can lead to spurious conclusions since not all time series properties are invariant to aggregation. The divergence from the actual causal relationship between two time series and the false relationship is known as temporal aggregation bias, commonly referred to as spurious causality. This effect has also been mentioned in [Breitung e Swanson](#page-28-2)  $(2002)$  and Götz et al.  $(2016)$ .

In a related study, [Renault et al.](#page-29-4) [\(1998\)](#page-29-4) investigates another reason for spurious causality, attributing it to the use of discrete data in causality analysis. The authors state that using discrete data overlooks events that occur within time intervals, essentially disregarding valuable information within a series of observations. They describe significant misleading causal effects in discrete time that do not emerge from the continuous-time generating process. For further discussions on using continuous models, see [Florens e Fougere](#page-28-8) [\(1996\)](#page-28-8) and [Comte e Renault](#page-28-9) [\(1996\)](#page-28-9).

[Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10) addresses the problem of spurious causality in a Mixed Frequency Vector Autoregressive (MF-VAR) setting by assuming that the low-frequency sampled process is not temporally aggregated but instead sampled at its true frequency. This approach effectively rules out any spurious causality by construction. The finite sample properties presented in [Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10) indicate that Granger testing within the MF-VAR approach exhibits higher asymptotic power for small differences in sampling frequencies, such as quarterly/monthly mixtures. Götz et al. [\(2016\)](#page-28-7) also proposed a causality test using MF-VAR, demonstrating that a Bayesian methodology improves the sensitivity of the Granger test when examining causality from a very high frequency to low frequency. This improvement is shown through tests of causality from daily to quarterly observations, with adjustments made to handle parameter proliferation. [Tank et al.](#page-29-10)  $(2019)$  conducts an investigation into Granger causality in mixed-frequency time series models, focusing specifically on the identifiability of the structural vector.

Following the methodology presented in [Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10), the series are stacked, resulting in an MF vector,  $\mathbf{X}(\tau)$ , described by equation

$$
\boldsymbol{X}(\tau)=\left[\boldsymbol{x}_H(\tau,1)',\ldots,\boldsymbol{x}_H(\tau,s)',\boldsymbol{x}_L(\tau)'\right]',
$$

where  $\tau \in \{1, \ldots, T\}$  is the low-frequency (LF) time index,  $x_L(\tau)$  represents the LF variable, and the set  $\{\boldsymbol{x}_H(\tau,1)',\ldots,\boldsymbol{x}_H(\tau,s)'\}$  consists of the high-frequency (HF) observations, with *s* being the frequency ratio between the HF and LF variables. For example, for each quarterly GDP observation, there are three observations of a monthly indicator. Specifically, if  $x_L(\tau)$  denotes a quarterly variable, then  $x_H(\tau, 1)$ represents the first month of each quarter,  $x_H(\tau, 2)$  the second month, and so on.

Under some standard assumptions, as discussed in [Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10), consider a bivariate case where the structural form for the MF-VAR includes the current lowfrequency variable and three lags of the high-frequency variable, given by

$$
\begin{bmatrix} 1 & 0 & 0 & 0 \ -d & 1 & 0 & 0 \ 0 & -d & 1 & 0 \ 0 & 0 & 0 & 1 \ \end{bmatrix} \begin{bmatrix} x_H(\tau,1) \\ x_H(\tau,2) \\ x_H(\tau,3) \\ x_L(\tau) \end{bmatrix} = \begin{bmatrix} 0 & 0 & d & c_1 \\ 0 & 0 & 0 & c_2 \\ 0 & 0 & 0 & c_3 \\ b_3 & b_2 & b_1 & a \end{bmatrix} \begin{bmatrix} x_H(\tau-1,1) \\ x_H(\tau-1,2) \\ x_H(\tau-1,3) \\ x_L(\tau-1) \end{bmatrix} + \begin{bmatrix} \xi_H(\tau,1) \\ \xi_H(\tau,2) \\ \xi_H(\tau,3) \\ \xi_L(\tau) \end{bmatrix}
$$

or  $\mathbf{N}\mathbf{X}(\tau) = \mathbf{M}\mathbf{X}(\tau-1) + \boldsymbol{\xi}(\tau)$ . It also assumed that  $x_H$  follows a AR(1) with coefficient *d*. The impact of lagged  $x_L$  on  $x_H$  is governed by  $c_1, c_2$ , and  $c_3$ .  $x_L$  follows a AR(1) with coefficient *a*. The impact of lagged  $x_H$  on  $x_L$  is governed by  $b_1, b_2,$ and  $b_3$ . Premultiply both sides of the structural form by

$$
\boldsymbol{N}^{-1} = \left[ \begin{array}{rrrr} 1 & 0 & 0 & 0 \\ d & 1 & 0 & 0 \\ d^2 & d & 1 & 0 \\ 0 & 0 & 0 & 1 \end{array} \right]
$$

to get the reduced form  $\mathbf{X}(\tau) = \mathbf{A}_1 \mathbf{X}(\tau - 1) + \boldsymbol{\epsilon}(\tau)$ , where

$$
\mathbf{A}_1 = \mathbf{N}^{-1} \mathbf{M} = \begin{bmatrix} 0 & 0 & d & \sum_{i=1}^1 d^{1-i} c_i \\ 0 & 0 & d^2 & \sum_{i=1}^2 d^{2-i} c_i \\ 0 & 0 & d^3 & \sum_{i=1}^3 d^{3-i} c_i \\ b_3 & b_2 & b_1 & a \end{bmatrix}
$$

and  $\boldsymbol{\epsilon}(\tau) = \boldsymbol{N}^{-1}\boldsymbol{\xi}(\tau)$ .

We can demonstrate that  $x_H$  does not cause  $x_L$  in the mixed frequency setting if and only if  $b_1 = b_2 = b_3 = 0$ . Non-causality from  $x_L$  to  $x_H$  involves  $c_1$ ,  $c_2$  and  $c_3$ , the AR(1) coefficient of  $x_H$ , as seen in the upper-right block of  $A_1$ .

Unlike [Ghysels et al.](#page-28-10)  $(2016)$  and Götz et al.  $(2016)$ , we assume that aggregated series are generated at a higher frequency, resulting in missing information once aggregated. Thus, in the following section, we illustrate how this assumption influences Granger causality analysis.

# <span id="page-11-0"></span>**3 Aggregation and Spurious Causality**

In this section, we build upon Granger's concept of causality as introduced in Section [2](#page-8-0) for two time series,  $y_t$  and  $x_t$ , and their aggregation. By illustrating the spurious causality effects using simple models, we can better understand its impact on more complex models. The series  $x_t$  and  $y_t$  have a causal relationship described as

<span id="page-11-1"></span>
$$
\begin{array}{rcl}\ny_t &=& \theta_1 y_{t-1} + \theta_2 x_{t-1} + v_{1,t} \\
x_t &=& \lambda_1 y_{t-1} + \lambda_2 x_{t-1} + v_{2,t}\n\end{array}, \quad v_{j,t} \sim N(0, \sigma^2), \text{ for } j = \{1, 2\}.
$$
\n(3.1)

with both  $y_t$  and  $x_t$  modeled after one lag of themselves and one of each other. The errors  $v_{i,t}$ ,  $j = \{1,2\}$ , are assumed to be independent and identically distributed. If the coefficients  $\theta_2 \neq 0$  or  $\lambda_1 \neq 0$ , this indicates that we are working with a scenario where lagged values of  $x_t$  are influential in present values of  $y_t$  or lagged values of  $y_t$ are influential in present values of *x<sup>t</sup>* , respectively. Therefore, we have four possible cases of Granger causality between  $x_t$  and  $y_t$ :

(i)  $x_t$  doesn't Granger cause  $y_t$  and  $y_t$  doesn't Granger cause  $x_t$ ,

 $\theta_2 = 0$  and  $\lambda_1 = 0$ ,  $x_t \nleftrightarrow y_t$  and  $y_t \nleftrightarrow x_t$ ,

(ii)  $x_t$  Granger cause  $y_t$  and  $y_t$  doesn't Granger cause  $x_t$ ,

$$
\theta_2 \neq 0
$$
 and  $\lambda_1 = 0$ ,  $x_t \to y_t$  and  $y_t \nleftrightarrow x_t$ 

(iii)  $x_t$  doesn't Granger cause  $y_t$  and  $y_t$  Granger cause  $x_t$ ,

$$
\theta_2 = 0
$$
 and  $\lambda_1 \neq 0$ ,  $x_t \nrightarrow y_t$  and  $y_t \rightarrow x_t$ 

(iv)  $x_t$  Granger cause  $y_t$  and  $y_t$  Granger cause  $x_t$ ,

$$
\theta_2 \neq 0
$$
 and  $\lambda_1 \neq 0$ ,  $x_t \to y_t$  and  $y_t \to x_t$ 

Now, let us focus on the effects of aggregation on the measurement of causality. For this, let  $Y_\tau$  and  $X_\tau$  be potentially aggregated time series composed of lagged values of  $y_t$  and  $x_t$ , respectively, where  $\tau = \{1, 2, \dots\}$  represents periods corresponding to  $t = \{s, 2s, \dots\}$ , with *s* being the frequency ratio between the high-frequency and the low-frequency variables. We can then analyze the effects of  $Y_{\tau-1}$  and  $X_{\tau-1}$  on *Y<sub>τ</sub>*, as expressed by

$$
Y_{\tau} = \Theta_1 Y_{\tau-1} + \Theta_2 X_{\tau-1} + u_t
$$

The coefficients  $\hat{\Theta}_1$  and  $\hat{\Theta}_2$  can be retrieved by least squares,

<span id="page-12-0"></span>
$$
\begin{bmatrix}\n\hat{\Theta}_1 \\
\hat{\Theta}_2\n\end{bmatrix} = det(\hat{M})^{-1} \begin{bmatrix}\n(X'_{\tau-1}X_{\tau-1})(Y'_{\tau-1}Y_{\tau}) - (X'_{\tau-1}Y_{\tau-1})(X'_{\tau-1}Y_{\tau}) \\
(Y'_{\tau-1}Y_{\tau-1})(X'_{\tau-1}Y_{\tau}) - (X'_{\tau-1}Y_{\tau-1})(Y'_{\tau-1}Y_{\tau})\n\end{bmatrix} (3.2)
$$

where  $det(\hat{M}) = (Y'_{\tau-1}Y_{\tau-1})(X'_{\tau-1}X_{\tau-1}) - (X'_{\tau-1}Y_{\tau-1})^2$ . Furthermore, under usual time series linear regression assumptions and assuming the errors are i.i.d., we have

<span id="page-12-1"></span>
$$
E\begin{bmatrix} \hat{\Theta}_1\\ \hat{\Theta}_2 \end{bmatrix} = \begin{bmatrix} \Theta_1\\ \Theta_2 \end{bmatrix} = det(M)^{-1} \begin{bmatrix} \pi_1\\ \pi_2 \end{bmatrix}, \tag{3.3}
$$

where  $\pi_1$  and  $\pi_2$  are related to the second part of the right-hand side of the equation  $(3.2).$  $(3.2).$ 

Considering the relationship between the variables  $x_t$  and  $y_t$  as depicted in equation [\(3.1\)](#page-11-1), our objective is to illustrate how spurious conclusions about the relationship between them can emerge in the aggregate setting, utilizing the series  $Y_\tau$  and  $X_{\tau}$ . The following four examples spotlight instances of what may be deemed spurious causality between these series. These examples help to describe scenarios in testing for Granger causality that might be prone to the spurious causality problem.

Notice that, for the next examples, the series are not subsampled but only aggregated, as subsampling does not cause spurious causality. To illustrate this, assume that model [\(3.1\)](#page-11-1) is sampled every two observations. In this scenario,  $\{y_{t-1}, x_{t-1}\}$ would not be observable, but  $\{y_{t-2}, x_{t-2}\}$  would be. Consequently, even if  $\theta_2 = 0$ and  $\lambda_1 \neq 0$ , the projection of  $y_t$  on  $\{y_{t-2}, x_{t-2}\}$  would result in  $E[\tilde{\theta}_1] = \theta_1^2$  and  $E[\tilde{\Theta}_2] = 0$ , thus not leading to spurious causality. This simplification allows us to derive the effects of spurious causality without loss of generality.

#### **Example 1**

In this initial example, we assume Case (i),  $\theta_2 = 0$  and  $\lambda_1 = 0$ , i.e.,  $x_t \nrightarrow y_t$  and  $y_t \nightharpoonup x_t$ . Thus,

$$
\begin{array}{rcl}\ny_t &=& \theta_1 y_{t-1} + v_{1,t} \\
x_t &=& \lambda_2 x_{t-1} + v_{2,t} \n\end{array}, \quad v_{j,t} \sim N(0, \sigma^2), \text{ for } j = \{1, 2\}
$$
\n(3.4)

with a simple aggregation pattern defined by

<span id="page-12-2"></span>
$$
Y_{\tau} = y_t + y_{t-1}, \quad Y_{\tau-1} = y_{t-1} + y_{t-2}, \quad X_{\tau-1} = x_{t-1} + x_{t-2}.
$$
 (3.5)

Intuitively, we consider a present value of the *y<sup>t</sup>* series and a further lagged value of the  $x_t$  series. Solving for  $\Theta_1$  and  $\Theta_2$  in equations [\(3.2\)](#page-12-0) and [\(3.3\)](#page-12-1), we obtain

$$
det(M) = 4\sigma_x^2 \sigma_y^2 + 4\theta_1 \sigma_x^2 \sigma_y^2 + 4\lambda_2 \sigma_x^2 \sigma_y^2 + 4\theta_1 \lambda_2 \sigma_x^2 \sigma_y^2, \pi_1 = 4\theta_1 \sigma_x^2 \sigma_y^2 + 2\sigma_x^2 \sigma_y^2 + 2\theta_1^2 \sigma_x^2 \sigma_y^2 + 4\theta_1 \lambda_2 \sigma_x^2 \sigma_y^2 + 2\lambda_2 \sigma_x^2 \sigma_y^2 + 2\theta_1^2 \lambda_2 \sigma_x^2 \sigma_y^2, \pi_2 = 0.
$$

The result shows that  $E[\hat{\Theta}_2]$  equals zero, indicating that past values of  $X_\tau$  have no influence on current values of  $Y<sub>\tau</sub>$ . This is similar to the relationship between  $x_t$  and  $y_t$ . Hence, the non-causality among the non-aggregated series results in no spurious causality when using the aggregated setting.

#### **Example 2**

Now, we assume Case (iii),  $\theta_2 = 0$  but  $\lambda_1 \neq 0$ , i.e.,  $y_t \to x_t$  but  $x_t \nrightarrow y_t$ , or

<span id="page-13-1"></span>
$$
y_t = \theta_1 y_{t-1} + v_{1,t}
$$
  
\n
$$
x_t = \lambda_1 y_{t-1} + \lambda_2 x_{t-1} + v_{2,t}, \quad v_{j,t} \sim N(0, \sigma^2), \text{ for } j = \{1, 2\}
$$
 (3.6)

with the same aggregation patter defined in the first example, see equation  $(3.5)$ , the relation between  $Y_\tau$  given  $Y_{\tau-1}$  and  $X_{\tau-1}$  is given by

<span id="page-13-2"></span>
$$
det(M) = \sigma_x^2 \sigma_y^2 (4 + 4\theta_1 + 4\lambda_2 + 4\theta_1 \lambda_2) - \lambda_1^2 \sigma_y^4 + \sigma_{xy} (-4^2 - 4\theta_1^2 - \theta_1^2 + 2\theta_1 \lambda_1 \sigma_y^2 - 4\lambda_2^2 - 2\lambda_2 \lambda_1 \sigma_y^2 - \lambda_2^2 - 2\theta_1 \lambda_2^2),
$$
  

$$
\pi_1 = \sigma_x^2 \sigma_y^2 (4\theta_1 + 2 + 2\theta_1^2 + 4\theta_1 \lambda_2 + 2\lambda_2 + 2\theta_1^2 \lambda_2) +
$$
  

$$
\sigma_{xy}^2 (-5\theta_1 - 4\theta_1^2 - 2 - \theta_1^3 - 2\theta_1 \lambda_2 - \lambda_2 - \theta_1^2 \lambda_2) +
$$
  

$$
\sigma_y^2 \sigma_{xy} (+2\theta_1 \lambda_1 + \lambda_1 + \theta_1^2 \lambda_1) +
$$
  

$$
+ \sigma_{xy}^2 (2\theta_1 \lambda_1 \sigma_y^2 + 1\lambda_1 \sigma_y^2 + \theta_1^2 \lambda_1 \sigma_y^2),
$$
  

$$
\pi_2 = \sigma_y^2 \sigma_{xy} (\theta_1 + 2\theta_1^2 + \theta_1^3 - 2\theta_1 \lambda_2 - \lambda_2 - \theta_1^2 \lambda_2).
$$
  

$$
+ \sigma_{xy}^4 (-2\theta_1 \lambda_1 - \lambda_1 - \theta_1^2 \lambda_1).
$$

Despite the absence of  $\theta_2$  in the model, the results indicate that  $\Theta_2$  is likely different from zero, demonstrating a spurious causality caused by aggregation. The aggregated time series reveals an influence of  $X<sub>\tau</sub>$  on the  $Y<sub>\tau</sub>$  series, even though such an influence is non-existent when the time series is non-aggregated.

We calculate potential values for the  $\Theta_2$  coefficient, assuming all  $\sigma_{xy}$ ,  $\sigma_y$ , and  $\sigma_x$  as functions of  $\lambda_1$ ,  $\lambda_2$ , and  $\theta_1$ , with  $\sigma_{v,j} = 1$  for  $j = 1, 2$ . Setting  $\theta_2 = 0$  and varying all other coefficients between -0.95 and 0.95 in increments of 0.01, Figure [3.1](#page-13-0) illustrates the frequency of the expected estimator. The chart shows that values of  $\Theta_2$  are distinct from zero in several cases, making the spurious causality evident. It is important to note that although the estimations of  $\Theta_2$  are evenly distributed around zero, this does not imply that its true value is expected to be zero. For instance, with  $\lambda_1 = \lambda_2 = \theta_1 = 0.3$ , the expected value for  $\Theta_2$  equals -0.068.



<span id="page-13-0"></span>Figure 3.1: Histogram of possible values for  $\Theta_2$  accordingly to model [\(3.6\)](#page-13-1), equation [\(3.3\)](#page-12-1), and aggregation [\(3.5\)](#page-12-2). With  $\theta_2 = 0$  and  $\lambda_1$ ,  $\lambda_2$ , and  $\theta_1$  between -0.95 and 0.95 in increments of 0.01.

#### <span id="page-14-1"></span>**Example 3**

Similarly to previous example, we assume Case (iii), where  $x_t \nrightarrow y_t$  and  $y_t \rightarrow x_t$ , as presented in the model [\(3.6\)](#page-13-1). However, we now assume that  $X<sub>\tau</sub>$  is non-aggregated, thus measuring Granger causality from a non-aggregated to an aggregated series, i.e.,

 $Y_{\tau} = y_t + y_{t-1}, \quad Y_{\tau-1} = y_{t-1} + y_{t-2}, \quad X_{\tau-1} = x_{t-2}$  (3.7)

With this aggregation pattern, our objective is to assess whether there are any spurious effects when only one of the series is aggregated. In this setup, the furthest lag of the  $x_t$  series coincides with the furthest lag of the  $y_t$  series. The results are given by

<span id="page-14-2"></span>
$$
det(M) = 2\sigma_x^2 \sigma_y^2 + 2\theta_1 \sigma_x^2 \sigma_y^2 - 3\theta_1 \sigma_{xy}^2 - \sigma_{xy}^2,
$$
  
\n
$$
\pi_1 = 2\theta_1 \sigma_x^2 \sigma_y^2 + \sigma_x^2 \sigma_y^2 + \theta_1^2 \sigma_x^2 \sigma_y^2 - \theta_1^3 \sigma_{xy}^2 - 2\theta_1^2 \sigma_{xy}^2 - \theta_1 \sigma_{xy}^2,
$$
  
\n
$$
\pi_2 = \theta_1^2 \sigma_y^2 \sigma_{xy} + \theta_1^3 \sigma_y^2 \sigma_{xy} - \theta_1 \sigma_y^2 \sigma_{xy} - \sigma_y^2 \sigma_{xy}.
$$

Notice that the expected value of  $\hat{\Theta}_2$  still deviates from zero, indicating a per-sistent issue of spurious causality. As shown in Figure [3.2,](#page-14-0) the distribution of  $\Theta_2$ appears more concentrated around zero compared to Figure [3.1.](#page-13-0)



Figure 3.2: Histogram of possible values for  $\Theta_2$  accordingly to model [\(3\)](#page-14-1), equation [\(3.3\)](#page-12-1), and aggregation [\(3.7\)](#page-13-2). With  $\theta_2 = 0$  and  $\lambda_1$ ,  $\lambda_2$ , and  $\theta_1$  between -0.95 and 0.95 in increments of 0.01.

One could argue that introducing a higher lag might reduce substantially the spurious effect. However, although this is true, it will not eliminate the spurious causality. To test if a higher lag value of  $x_t$  is sufficient to reduce spurious effects considerably, we replicate the calculations and simulations mentioned earlier, this time with  $X_{\tau-1}$  being non-aggregated and equal to  $x_{t-5}$ .

<span id="page-14-0"></span>
$$
Y_{\tau} = y_t + y_{t-1}, \quad Y_{\tau-1} = y_{t-1} + y_{t-2}, \quad X_{\tau-1} = x_{t-5}
$$
 (3.8)

with results given by

<span id="page-14-3"></span>
$$
det(M) = 2\sigma_x^2 \sigma_y^2 + 2\theta_1 \sigma_x^2 \sigma_y^2 - 3\theta_1 \sigma_{xy}^2 - \sigma_{xy}^2,
$$
  
\n
$$
\pi_1 = 2\theta_1 \sigma_x^2 \sigma_y^2 + \sigma_x^2 \sigma_y^2 + \theta_1^2 \sigma_x^2 \sigma_y^2 - \theta_1^3 \sigma_{xy}^2 - 2\theta_1^2 \sigma_{xy}^2 - \theta_1 \sigma_{xy}^2,
$$
  
\n
$$
\pi_2 = \theta_1^5 \sigma_y^2 \sigma_{xy} + \theta_1^6 \sigma_y^2 \sigma_{xy} - \theta_1^4 \sigma_y^2 \sigma_{xy} - \theta_1^3 \sigma_y^2 \sigma_{xy}.
$$

The distribution of  $\Theta_2$  seen in Figure [3.3](#page-15-0) is now closer to zero than in the previous example, but it is not exactly zero. Since it differs from zero, the rejection rate of  $H_0: \hat{\Theta}_2 = 0$  will tend to 1 as  $n \to \infty$ .



<span id="page-15-0"></span>Figure 3.3: Histogram of possible values for  $\Theta_2$  accordingly to model [\(3\)](#page-14-1), equation [\(3.3\)](#page-12-1), and aggregation [\(3.8\)](#page-14-2). With  $\theta_2 = 0$  and  $\lambda_1$ ,  $\lambda_2$ , and  $\theta_1$  between -0.95 and 0.95 in increments of 0.01.

#### **Example 4**

Finally, we examine the same scenario, Case (iii),  $x_t \nrightarrow y_t, y_t \rightarrow x_t$ , model [\(3.6\)](#page-13-1), but this time, we reverse the aggregation, with  $Y_\tau$  and  $Y_{\tau-1}$  being non-aggregated and  $X_{\tau}$  being aggregated, i.e.,

$$
Y_{\tau} = y_t, \quad Y_{\tau-1} = y_{t-1}, \quad X_{\tau-1} = x_{t-1} + x_{t-2}.
$$
 (3.9)

This allows us to investigate the causality direction from an aggregated series to a non-aggregated one. The results are given by

$$
det(M) = 2\sigma_x^2 \sigma_y^2 - \sigma_{xy}^2 - 2\theta_1 \sigma_{xy}^2 - \theta_1^2 \sigma_{xy}^2 + 2\lambda_1 \sigma_y^2 \sigma_{xy} + 2\lambda_2 \sigma_x^2 \sigma_y^2, \pi_1 = 2\theta_1 \sigma_x^2 \sigma_y^2 - \theta_1 \sigma_{xy}^2 - 2\theta_1^2 \sigma_{xy}^2 - \theta_1^3 \sigma_{xy}^2 + 2\theta_1 \lambda_1 \sigma_y^2 \sigma_{xy} + 2\theta_1 \lambda_2 \sigma_x^2 \sigma_y^2, \pi_2 = 0.
$$

Notice that the result for  $\Theta_2$  is now zero, indicating that there are no issues related to spurious causality when the  $x_t$  series is not aggregated, contrary to Example 3. In summary, testing Granger causality with an aggregated series, such as  $Y_{\tau} = y_t + y_{t-1}$ , and a non-aggregated series, such as  $x_t$ , allows for the correct measurement of causality from  $y_t$  to  $x_t$ .

**REMARK:** We derived the above results by implementing an algorithm tailored to this purpose. Written in  $R$  and adjusted to model  $(3.1)$ , the algorithm efficiently solves the symbolic expressions. In the Appendix, we present the results for the GDP aggregation, which are explained in Section [6.](#page-20-0)

### <span id="page-17-0"></span>**4 Sims-causality for Mixed-Frequencies**

According to [Chamberlain](#page-28-11) [\(1982\)](#page-28-11), Granger causality is defined such that if  $x_t$  does not cause *y<sup>t</sup>* then

<span id="page-17-1"></span>(G)  $y_t$  is independent of  $x_{t-1}, x_{t-2}, \ldots$ , conditional on  $y_{t-1}, y_{t-2}, \ldots, \forall t$  (4.1)

in contrast, [Sims](#page-29-6) [\(1972\)](#page-29-6) proposed a distinct measure of causality, or more precisely, the measurement of strict exogeneity, where if  $x_t$  does not cause  $y_t$ , then the regression of  $x_t$  against leads and lags of  $y_t$  will result in null coefficients for the lead variables. The Sims's causality definition by [Chamberlain](#page-28-11) [\(1982\)](#page-28-11) is

<span id="page-17-2"></span>(S)  $x_t$  is independent of  $y_{t+1}, y_{t+1}, \ldots$ , conditional on  $y_t, y_{t-1}, \ldots, \forall t$  (4.2)

[Chamberlain](#page-28-11) [\(1982\)](#page-28-11) argue that non-causality is a more stringent criterion than strict exogeneity, suggesting that equation  $(4.1)$  implies equation  $(4.2)$  while the inverse is not necessarily true. This result is similar to the one presented in [Florens](#page-28-12) [e Mouchart](#page-28-12) [\(1982\)](#page-28-12), where is also pointed that both conditions, equations [\(4.1\)](#page-17-1) and [\(4.2\)](#page-17-2), became equivalent when *x<sup>t</sup>* and *y<sup>t</sup>* are Gaussian processes. [Kuersteiner](#page-29-11) [\(2010b\)](#page-29-11) shows that the equivalence between Sims and Granger causality no longer holds when additional covariates are included in the analysis. The author demonstrates that when considering three time series  $z_t$ ,  $y_t$ , and  $x_t$  in a causality relation where  $z_t \to x_t$ ,  $z_t \to y_t$ , and  $x_t$  is not directly causing  $y_t$ , using the Granger approach we will conclude that  $x_t \nightharpoonup y_t$ . However, using the Sims approach, we would wrongly conclude that  $x_t \to y_t$ , demonstrating that the Sims approach is not suited for multivariate systems.

We know from Example 4, Section [3,](#page-11-0) that temporal aggregation does not compromise the measurement of causality between the aggregated variable and the regular variable. The same holds for Sims causality, which can be measured using

<span id="page-17-3"></span>
$$
Y_{\tau} = \sum_{j=-p}^{p} \gamma_{t-j} x_{t-j} + v_t
$$
\n(4.3)

where  $\tau = \{1, 2, \ldots\}$  represents periods corresponding to  $t = \{s, 2s, \ldots\}$ , with *s* been the frequency ratio between the high-frequency and the low-frequency variables. If any coefficients of  $\{x_{t+1}, \ldots, x_{t+p}\}$ , or the conjunction of all coefficients, are significant then we say that  $Y_\tau$  Sims-cause  $x_t$ .

To test if  $x_t$  causes  $Y_\tau$ , we propose a modification to the Sims test that avoids any possible overlap between future values of  $Y_\tau$  and the present value of  $x_t$ . This modified Sims test can be described by

<span id="page-18-0"></span>
$$
x_t = \sum_{j=0}^{\infty} \gamma_{t-j} Y_{\tau-j} + \sum_{j=1}^{p} \gamma_{t+j} Y_{\tau+j} I(j) + v_t
$$
\n(4.4)

where  $I(j)$  is a indicator function that is equal to zero when the aggregated variable  $Y_{\tau+j} = y_{(\tau+j)\cdot s} + y_{(\tau+j)\cdot s-1} + \cdots + y_{(\tau+j)\cdot s-k}$  contains any element that occurred at same time or before *x<sup>t</sup>* .

**Theorem 1.** *Given the conditions of [Sims](#page-29-6) [\(1972\)](#page-29-6) and equations [\(4.3\)](#page-17-3)-[\(4.4\)](#page-18-0), if*  $Y_{\tau} \to x_t$  or  $Y_{\tau} \nrightarrow x_t$ , then  $y_t \to x_t$  or  $y_t \nrightarrow x_t$ , respectively. Also if  $x_t \to Y_{\tau}$  or  $x_t \nightharpoonup Y_\tau$  *then*  $x_t \rightarrow y_t$  *or*  $x_t \nightharpoonup y_t$ *, respectively.* 

*Sketch.* The proof of equation [\(4.3\)](#page-17-3) is straightforward and is therefore omitted. For equation [\(4.4\)](#page-18-0), the proof is also straightforward when both  $x_t \nrightarrow y_t$  and  $y_t \nrightarrow x_t$ . However, when  $x_t \nrightarrow y_t$  but  $y_t \rightarrow x_t$ , we have the system

$$
y_t = \theta_1 y_{t-1} + v_{1,t},
$$
  
\n
$$
x_t = \lambda_1 y_{t-1} + \lambda_2 x_{t-1} + v_{2,t}.
$$

Note that  $x_t$  can be expressed as a function solely of past values of  $y_t$ , specifically as

$$
x_t = \sum_{j=1}^{\infty} \lambda_1 \lambda_2^{j-1} y_{t-j} + v_{2,t} + \sum_{j=1}^{\infty} \lambda_2^j v_{2,t-j},
$$

or, more simply, as

$$
x_t = \sum_{j=1}^{\infty} \tilde{\lambda}_j y_{t-j} + \tilde{v}_{2,t}.
$$

Substituting this into equation  $(4.4)$ , we obtain

$$
\left(\sum_{j=1}^{\infty} \tilde{\lambda}_1 y_{t-j} + \tilde{v}_{2,t}\right) = \sum_{j=0}^{\infty} \gamma_{t-j} Y_{\tau-j} + \sum_{j=1}^{p} \gamma_{t+j} Y_{\tau+j} I(j) + v_t,
$$

which simplifies to

$$
(1+\tilde{\lambda}_jL+\cdots+\tilde{\lambda}_jL^k)y_t=\sum_{j=0}^{\infty}\gamma_{t-j}Y_{\tau-j}+\sum_{j=1}^p\gamma_{t+j}Y_{\tau+j}I(j)+\tilde{v}_t.
$$

Given that  $Y_\tau = (1 + a_1 L + \cdots + a_k L^k) y_t$ , where the  $a_i, i = 1, \ldots, k$ , terms are the temporal aggregation weights and the fact that *y<sup>t</sup>* can be also be written in terms of its past, the infinite summation on the right-hand side of [\(4.4\)](#page-18-0) can fully explain the left-hand side. Thus leading to the  $\gamma_{t+j}$ ,  $j=1,\ldots,p$ , coefficients to be zero.  $\Box$ 

The theorem above states that, assuming  $Y_\tau$  to be a potentially aggregated time series composed of lagged values of *y<sup>t</sup>* , any conclusion made regarding the causality relationship between  $Y_\tau$  and  $x_t$  will also be drawn between the non-aggregated  $y_t$  and *xt* . In other words, any causality conclusion made when considering an aggregated quarterly variable, such as GDP, and any other monthly indicator can be extended to the relationship between the same monthly variable and GDP in its non-aggregated state. In Section [6,](#page-20-0) we will present, by simulation, how the implementation of the proposed Sims test reflects the behavior presented by the theorem.

# <span id="page-19-0"></span>**5 Technical Procedures**

Considering two time series  $x_t$  and  $y_t$  and their aggregated versions,  $X_\tau$  and  $Y_\tau$ , composed of lagged values of  $x_t$  and  $y_t$ , respectively, we will assess the causality cases described as  $x_t \to Y_\tau$ ,  $Y_\tau \to x_t$ ,  $X_\tau \to Y_\tau$ , and  $Y_\tau \to X_\tau$ . In these cases, we will evaluate the performance of the tests mentioned in previous sections by implementing each test and comparing their sensitivity to spurious effects.

To analyze the Granger non-causality testing behavior, we follow the structure shown in equation  $(2.1)$ . By defining the order as one, we choose to consider a single lag for each series. This corresponds to using the same equation structure presented in equation [\(3.1\)](#page-11-1) for each  $x_t$  and  $y_t$  series. We refer to this implementation as  $G_{yx}(1)$ . Note that this test will be only used to test the aggregation explained in Example 4, equation [\(3.9\)](#page-14-3).

Two tests were implemented considering the Sims causality testing methodology presented in Chapter [4.](#page-17-0) The first test is the Sims test with a fixed order, represented as  $S_{xy}(1)$ , where we consider the number of lags and leads to be equal to one. For the  $x_t \to Y_\tau$  scenario, we follow equation [\(4.3\)](#page-17-3), and for the  $Y_\tau \to x_t$  scenario, we follow equation  $(4.4)$ . In the second test, we determine the maximum number of lags for each series using the function  $0.5 \cdot (n/s)^{(1/3)}$ . We refer to this test as  $S_{xy}(k)$ , with *k* being the test order represented by the lag function. In both tests, the number of lead observations is determined by the direction of causality being tested. For the  $x_t \to Y_\tau$  scenario, we consider the lead to be 1. For the  $Y_\tau \to x_t$  scenario, we consider the lead to be 2, skipping the first lead of  $Y_\tau$ .

Another causality test is conducted on both  $y_t$  and  $x_t$  in their original frequencies and without aggregation. This test uses the Granger methodology to serve as the benchmark scenario for causality testing between the two series. We refer to this implementation as the *Non-Agg* testing approach. In contrast, the *Agg* implementation involves aggregating the higher frequency series to match the lower frequency series before performing the Granger test in both directions:  $X_{\tau} \to Y_{\tau}$ and  $Y_\tau \to X_\tau$ . However, as discussed in Section [3,](#page-11-0) this method introduces spurious causality effects. The objective of this implementation is to investigate these effects further and compare them with the results obtained using the proposed tests.

### <span id="page-20-0"></span>**6 Finite sample properties**

In this chapter, we aim to simulate a fundamental relationship in Economics: the interaction between GDP and monthly indicators. GDP observations are reported quarterly, without overlap months, and are obtained by summing the GDP values of the preceding three months, i.e.,  $Y_\tau = \bar{y}_t + \bar{y}_{t-1} + \bar{y}_{t-2}$  where *m* denotes monthly observations. As shown in [Taufemback](#page-30-3) [\(2023\)](#page-30-3), the difference between the current quarter's GDP,  $Y_{\tau}$ , and the previous quarter,  $Y_{\tau-1}$ , reflects changes or innovations within that quarter. Equation  $(6.1)$  illustrates how monthly GDP innovations are structured following a quarterly difference, under the assumption that  $\bar{y}_t = \sum_{j=0}^{\infty} y_{t-j}$ , where  $y_{t-j}$  are the monthly innovations and  $\bar{y}_t \sim I(1)$ , with  $\Delta_q$ denoting the difference between consecutive quarters.

<span id="page-20-1"></span>
$$
\Delta_q Y_\tau = \bar{y}_t + \bar{y}_{t-1} + \bar{y}_{t-2} - (\bar{y}_{t-3} + \bar{y}_{t-4} + \bar{y}_{t-5})
$$
\n
$$
= \sum_{j=0}^\infty y_{t-j} + \sum_{j=1}^\infty y_{t-j} + \sum_{j=2}^\infty y_{t-j} - \sum_{j=3}^\infty y_{t-j} - \sum_{j=4}^\infty y_{t-j} - \sum_{j=5}^\infty y_{t-j}
$$
\n
$$
= y_t + 2y_{t-1} + 3y_{t-2} + 2y_{t-3} + y_{t-4}.
$$
\n(6.1)

Now, let the following models be defined by

 $\text{Model 1: } y_t = 0.6y_{t-1} + u_{1t},$  $x_t = 0.5x_{t-1} - 0.2x_{t-2} + u_{2t}$ Model 2:  $y_t = 0.25y_{t-1} + 0.5x_{t-1} + u_{2t}$  $x_t = 0.65x_{t-1} + u_{1t}$ Model 3:  $y_t = 0.65y_{t-1} + u_{1t}$  $x_t = 0.25y_{t-1} + 0.5x_{t-1} + u_{2t}$ Model 4:  $y_t = 0.65y_{t-1} + 0.15x_{t-1} + u_{1t}$  $x_t = -0.25x_{t-1} + 0.55y_{t-1} + u_{2t}$ Model 5:  $y_t = 0.5y_{t-1} - 0.2y_{t-2} + 0.4x_{t-1} - 0.1x_{t-2} + 0.15x_{t-3} - 0.05x_{t-4} + u_{1t}$  $x_t = 0.65x_{t-1} + u_{2t}$ Model 6: *y<sup>t</sup>* = 0*.*65*yt*−<sup>1</sup> + *u*1*<sup>t</sup>*

Model 6: 
$$
y_t = 0.65y_{t-1} + u_{1t}
$$
,  
\n $x_t = 0.5x_{t-1} - 0.2x_{t-2} + 0.4y_{t-1} - 0.1y_{t-2} + 0.15y_{t-3} - 0.05y_{t-4} + u_{2t}$ .

The six models presented above were designed to illustrate the different causality effects discussed in Section [3.](#page-11-0) For each series, the errors  $u_1$  and  $u_2$  were generated by

sampling from a i.i.d. normal distribution, and the aggregation of series  $y_t$  follows Equation  $(6.1)$ . In Model 1, both  $x_t$  and  $y_t$  consist solely of their lags, indicating the absence of a causal relationship between the series. Model 2 introduces a scenario where  $y_t$  includes one lag of  $x_t$ , while  $x_t$  depends only on its past values, suggesting  $x_t$  causes  $y_t$  but not vice versa. Conversely, Model 3 demonstrates  $y_t$  causing  $x_t$ while  $x_t$  does not cause  $y_t$ . Model 4 incorporates lagged values from both series, indicating a bidirectional causality scenario. In Model 5,  $y_t$  includes lags from  $x_t$ lags up to *t* − 4 and contains only one lag of itself. Finally, Model 6 mirrors the structure of Model 5 but with the roles of  $x_t$  and  $y_t$  reversed.

Our findings, as shown in Tables [6.1](#page-23-0) and [6.2,](#page-24-0) indicate that both the  $S_{xy}(k)$ and *Sxy*(1) tests maintain rejection rates closer to the *Non-Agg* test, demonstrating correct sensitivity in detecting causal relationships between variables. Conversely, the *Agg* test results suggest spurious causality, as expected. The rejection rate of the  $G_{yx}(1)$  test gradually increases to one when the alternative hypothesis is true, particularly for models 3, 4, and 6 in Table [6.2.](#page-24-0) However, this increase is consistently slower than that observed for the  $S_{xy}(k)$  and  $S_{xy}(1)$  tests. Notably, in cases where the *Agg* tests diverge, both the  $S_{xy}(k)$  and  $S_{xy}(1)$  tests maintain rejection rates around 0.05.

Analyzing each model individually, for Model 1 in Table [6.1,](#page-23-0) the *Non-Agg* test rejection rate presents values around 0.05. The *Agg* test rejection rate is slightly over-sized for  $n \in \{80, 120\}$  but stabilizes around 0.05 for  $n \in \{160, 400\}$ . Both Sims tests,  $S_{xy}(1)$  and  $S_{xy}(k)$ , are slightly under-sized. In Table [6.2,](#page-24-0) the *Non-Agg* test is under-sized for all sample sizes of *n*, while the *Agg* test maintains a rejection rate around 0.05. The  $G_{yx}(1)$  test is also under-sized but remains close to 0.05. Both  $S_{xy}(1)$  and  $S_{xy}(k)$  tests consistently vary around 0.05, which is desirable.

For Model 2, as shown in Table [6.1,](#page-23-0) the *Non-Agg* test consistently exhibits a 100% rejection rate across all sample sizes of *n*. Similarly, the rejection rate for the *Agg* test approaches 100% as *n* increases. Both the  $S_{xy}(1)$  and  $S_{xy}(k)$  tests also converge to 100% rejection rates, with  $S_{xy}(1)$  showing a higher rejection rate than  $S_{xy}(k)$ . Table [6.2](#page-24-0) shows that the *Non-Agg* test consistently maintains rejection rates around 0.05 across different sample sizes. However, the rejection rate of the *Agg* test diverges. The  $G_{yx}(1)$  test, along with the  $S_{xy}(1)$  and  $S_{xy}(k)$  tests, shows rejection rates close to 0.05.

For Model 3, as shown in Table [6.1,](#page-23-0) the *Non-Agg* test maintains a stable rejection rate around 0.05 across all sample sizes. In contrast, the *Agg* test exhibits rejection rates that diverge from 0.05 as *n* increases. The  $S_{xy}(k)$  test also maintains rejection rates close to 0.05, although they are slightly elevated for  $n \in \{120, 160\}$ . Conversely,  $S_{xy}(1)$  shows values closer to the expected 0.05. In Table [6.2,](#page-24-0) the *Non-Agg* test exhibits a 100% rejection rate. Similarly, the *Agg* test approaches a 100% rejection rate as *n* increases. The  $G_{yx}(1)$  test shows a gradual increase in rejection rates with increasing *n*. Both Sims tests,  $S_{xy}(k)$  and  $S_{xy}(1)$ , also demonstrate increasing rejection rates as *n* grows, converging towards 100

For Model 4, as shown in Table [6.1,](#page-23-0) both the *Non-Agg* and *Agg* tests exhibit rejection rates approaching 100% as *n* increases. Similarly, both Sims tests show increasing rejection rates with increasing *n*, approaching 100%. In Table [6.2,](#page-24-0) the *Non-Agg* test consistently shows a 100% rejection rate across all sample sizes. Likewise, the *Agg* test demonstrates a rejection rate that converges to 100% as *n* increases. The Granger  $G_{yx}(1)$  test exhibits a very low rejection rate, with values increasing slowly as *n* grows. Both Sims tests,  $S_{xy}(k)$  and  $S_{xy}(1)$ , also show increasing rejection rates as *n* increases, nearing  $100\%$  for  $n = 400$ .

For Model 5, as presented in Table [6.1,](#page-23-0) both the *Non-Agg* and *Agg* tests exhibit a 100% rejection rate. Similarly, both Sims tests show rejection rates that converge to 100% as *n* increases. In Table [6.2,](#page-24-0) the *Non-Agg* test maintains a rejection rate close to 0.05 across different sample sizes. Conversely, the *Agg* test demonstrates high rejection rates that increase significantly as *n* grows. The Granger  $G_{yx}(1)$  test initially shows a slightly inflated rejection rate for  $n = 80$ , which converges towards 0.05 as *n* increases. Both Sims tests,  $S_{xy}(k)$  and  $S_{xy}(1)$ , exhibit rejection rates close to 0.05 across sample sizes.

For Model 6, as shown in Table [6.1,](#page-23-0) the rejection rate of the *Non-Agg* test is slightly elevated for  $n = 80$ , but remains around 0.05 for other values of *n*. The *Agg* test exhibits consistently high rejection rates that increase with *n*. Both the  $S_{xy}(k)$ and  $S_{xy}(1)$  tests show rejection rates slightly below 0.05, but close to the expected level. In Table [6.2,](#page-24-0) the *Non-Agg* test demonstrates a 100% rejection rate across all sample sizes. Similarly, the *Agg* test shows rejection rates close to 100%. The Granger  $G_{yx}(1)$  test exhibits a rejection rate that increases slowly as *n* increases. Both Sims tests,  $S_{xy}(k)$  and  $S_{xy}(1)$ , initially show low rejection rates for  $n = 80$ , but these rates converge towards 100% as *n* increases.

	Model 1 - $x_t \nrightarrow y_t, y_t \nrightarrow x_t$								
$n_{agg}$	$Non-Agg$	Agg	$S_{xy}(k)$	$S_{xy}(1)$					
80	0.055	0.060	0.037	0.043					
120	0.041	0.063	0.036	0.035					
160	0.050	0.059	0.046	0.044					
400	0.060	0.049	0.038	0.035					
Model 2 - $x_t \rightarrow y_t, y_t \nrightarrow x_t$									
$n_{agg}$	$Non-Agg$	Agg	$S_{xy}(k)$	$S_{xy}(1)$					
80	0.998	0.825	0.685	0.694					
120	1.000	0.955	0.861	0.868					
160	1.000	0.982	0.925	0.937					
400	1.000	1.000	1.000	1.000					
$\mathbf{Model\ 3}\text{ - }\mathbf{x_{t}}\nrightarrow \mathbf{y_{t}},\mathbf{y_{t}}\rightarrow \mathbf{x_{t}}$									
$n_{agg}$	$Non-Agg$	Agg	$S_{xy}(k)$	$S_{xy}(1)$					
80	0.049	0.210	0.065	0.067					
120	0.043	0.211	0.087	0.083					
160	0.047	0.266	0.072	0.069					
400	0.053	0.498	0.063	0.069					
	$\label{model4} \begin{aligned} \text{Model 4 - } x_t \rightarrow y_t, y_t \rightarrow x_t \end{aligned}$								
$n_{agg}$ 80	$Non-Agg$ 0.859	Agg 0.765	$S_{xy}(k)$ 0.294	$S_{xy}(1)$ 0.307					
120	0.941	0.856	0.361	0.370					
160	0.981	0.947	0.461	0.473					
400	1.000	1.000	0.832	0.839					
	Model 5 - $x_t \rightarrow y_t, y_t \nrightarrow x_t$								
	$Non-Agg$								
$n_{agg}$ 80	1.000	Agg 0.987	$S_{xy}(k)$ 0.762	$S_{xy}(1)$ 0.775					
120	1.000	0.999	0.911	0.910					
160	1.000	1.000	0.966	0.967					
400	1.000	1.000	1.000	1.000					
	Model 6 - $x_t \nrightarrow y_t, y_t \rightarrow x_t$								
$n_{agg}$	$Non-Agg$	Agg	$S_{xy}(k)$	$S_{xy}(1)$					
80	0.065	0.264	0.048	0.052					
120	0.052	0.356	0.042	0.044					
160	0.058	0.407	0.036	0.037					

<span id="page-23-0"></span>Table 6.1: Average simulated rejections for  $H_0: x_t \nrightarrow y_t$  with  $\alpha = 0.05$ .  $\mathbf{Model} \; \mathbf{1} \text{ - } \; \mathbf{x_t} \nrightarrow \mathbf{y_t}, \mathbf{y_t} \nrightarrow \mathbf{x_t}$ 

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**Note:** Results are obtained considering the aggregation presented in equation [\(6.1\)](#page-20-1), models from Section [6,](#page-20-0) and the procedures presented in Section [5.](#page-19-0)

Model 1 - $x_t \nrightarrow y_t, y_t \nrightarrow x_t$									
$n_{agg}$	$Non-Agg$	Agg	$G_{yx}(1)$	$S_{yx}(k)$	$S_{yx}(f)$				
80	0.049	0.051	0.052	0.052	0.049				
120	0.042	0.051	0.037	0.060	0.058				
160	0.046	0.049	0.043	0.055	0.051				
400	0.036	0.043	0.039	0.041	0.042				
Model 2 - $x_t \rightarrow y_t, y_t \nrightarrow x_t$									
$n_{aqq}$	$Non-Agg$	Agg	$G_{yx}(1)$	$S_{yx}(k)$	$S_{yx}(f)$				
80	0.055	0.154	0.052	0.047	0.044				
120	0.053	0.212	0.045	0.043	0.049				
160	0.048	0.266	0.046	0.058	0.048				
400	0.067	0.515	0.046	0.042	0.037				
	Model 3 - $x_t \nrightarrow y_t, y_t \rightarrow x_t$								
$n_{agg}$	$Non-Agg$	Agg	$G_{yx}(1)$	$S_{yx}(k)$	$S_{yx}(f)$				
80	0.998	0.834	0.092	0.341	0.345				
120	1.000	0.946	0.086	0.456	0.464				
160	1.000	0.985	0.106	0.624	0.624				
400	1.000	1.000	0.213	0.949	0.947				
$\label{model4} \begin{aligned} \text{Model 4 - } x_t \rightarrow y_t, y_t \rightarrow x_t \end{aligned}$									
$n_{agg}$ 80	$Non-Agg$ 0.999	Agg 0.947	$G_{yx}(1)$ 0.067	$S_{yx}(k)$ 0.344	$S_{yx}(f)$ 0.353				
120	1.000	0.983	0.061	0.481	0.496				
160	1.000	0.998	0.069	0.589	0.599				
400	1.000	1.000	0.116	0.938	0.943				
	Model 5 - $x_t \rightarrow y_t, y_t \nrightarrow x_t$								
$n_{agg}$	$Non-Agg$	Agg	$G_{yx}(1)$	$S_{yx}(k)$	$S_{yx}(f)$				
80 120	0.050 0.062	0.254 0.342	0.063 0.055	0.047 0.047	0.052 0.069				
160	0.048	0.432	0.058	0.052	0.058				
400	0.041	0.818	0.055	0.054	0.041				
	Model 6 - $x_t \nrightarrow y_t, y_t \rightarrow x_t$								
$n_{agg}$	$Non-Agg$	Agg	$G_{yx}(1)$	$S_{yx}(k)$					
80	1.000	0.982	0.045	0.636	$S_{yx}(f)$ 0.621				
120	1.000	0.999	0.057	0.831	0.809				
160 400	1.000 1.000	1.000 1.000	0.066 0.098	0.910 1.000	0.899 1.000				

<span id="page-24-0"></span>Table 6.2: Average simulated rejections for  $H_0: y_t \nightharpoonup x_t$  with  $\alpha = 0.05$ .

**Note:** Results are obtained considering the aggregation presented in equation [\(6.1\)](#page-20-1), models from Section [6,](#page-20-0) and the procedures presented in Section [5.](#page-19-0)

# <span id="page-25-0"></span>**7 Empirical analysis**

In this section, we apply the tests described in Section [5](#page-19-0) to U.S. macroeconomic series data within a bivariate framework. We consider the same set of variables as in [Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10), specifically U.S. inflation (CPI), monthly crude oil price fluctuations (OIL), and quarterly real GDP growth. Additionally, we include other U.S. indicators utilized in [Stock e Watson](#page-29-12) [\(1989\)](#page-29-12) and [Mariano e Murasawa](#page-29-13) [\(2003\)](#page-29-13), such as income per person (INC), industrial production (IP), industries sales (IS), and total employment (EMP). The dataset spans from January 1959 to December 2019, with all data publicly available and sourced from [research.stlouisfed.org.](https://research.stlouisfed.org/econ/mccracken/fred-databases/) Each series was made stationary following the guidelines outlined in [McCracken e Ng](#page-29-14) [\(2016\)](#page-29-14). Table [7.1](#page-25-1) presents the p-values for each test, along with the tested causality direction.

<span id="page-25-1"></span>

Test direction	Agg	$G_{yx}(1)$	$S_{xy}(k)$	$S_{xy}(1)$
OIL $\#$ GDP	0.0586		0.2771	0.3455
$GDP \nrightarrow OIL$	0.7002	0.3215	0.1483	0.3986
$CPI \nrightarrow GDP$	0.0000		0.5358	0.4368
$GDP \nrightarrow CPI$	0.0163	0.6380	0.1009	0.5459
INC $\rightarrow$ GDP	0.0000		0.0256	0.0132
$GDP \nrightarrow \text{INC}$	0.0000	0.0002	0.0000	0.0000
IP $\rightarrow$ GDP	0.0000		0.4341	0.9833
$GDP \nrightarrow IP$	0.1321	0.0081	0.0013	0.0156
$EMP \nrightarrow GDP$	0.0000		0.3490	0.2431
$GDP \nrightarrow EMP$	0.07609	0.0000	0.0269	0.0001
IS $\rightarrow$ GDP	0.0000		0.0077	0.0128
$GDP \nrightarrow IS$	0.0136	0.0000	0.0936	0.0001

Table 7.1: Empirical results

**Note:** All p-values highlighted in bold are inferior to the 5% rejection level.

The  $S_{xy}(k)$  test and the  $S_{xy}(1)$  test both fail to detect causality from OIL, CPI, IP, and EMP towards GDP. However, they do identify causality from GDP to INC, IP, EMP, and IS. The only discrepancy between the tests occurs in detecting causality from GDP to IS, where the  $S_{xy}(k)$  test only rejects at 10%. Neither test rejects the hypothesis of non-causality in either direction when examining GDP with OIL and CPI. The findings of [Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10) indeed indicate no causality between GDP and OIL, but they detect causality to and from GDP with respect to CPI.

The  $G_{yx}(1)$  test consistently detects causality from GDP to the monthly indicators in most scenarios despite its lower rejection power. However, similar to the  $S_{xy}(k)$  and  $S_{xy}(1)$  tests, it does not reject the hypothesis of non-causality when testing with CPI and OIL. This reinforces our findings and contrasts with those of [Ghysels et al.](#page-28-10) [\(2016\)](#page-28-10) regarding CPI.

As previously discussed, temporal aggregation can lead to spurious causality effects, which might explain why the *Agg* testing sometimes detects causality when other tests do not.

### <span id="page-27-0"></span>**8 Conclusion**

Aggregating high-frequency variables can result in missing information and potentially misleading conclusions when testing causality between time series. In this study, we propose a modification to the Sims causality test designed to address this issue. Specifically, the modification allows for detecting causality when one series is aggregated while the other remains not aggregated, thereby avoiding spurious conclusions.

In practice, it is challenging to draw conclusions about the causal relationship between economic variables solely based on a bivariate time series model. The structure of the relationship can only be accurately derived by including all relevant variables in the model. Consequently, since many economic variables interact and are important, high-dimensional time series model-building is necessary. As discussed in Lütkepohl [\(1982\)](#page-29-15), a low-dimensional sub-process may not fully capture the dynamics of a higher-dimensional system. Thus, even if  $x_t$  does not cause  $Y_\tau$ in the bivariate context,  $Y_\tau$  could still respond to changes in  $x_t$  within a broader, multivariate framework.

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# **Appendix**

The set of equations bellow represent the covariance between values of the  $y_t$  and  $x_t$ series considering a different lag *h*, for any  $h \in [1, 10]$ , for model [\(3.1\)](#page-11-1). We use  $\gamma_y(h)$ to represent a covariance between a present value  $y_t$  and  $y_{t-h}$ . We utilize  $\psi_{yx}(h)$ to represent a covariance between a present value  $y_t$  and  $x_{t-h}$ . We use  $\phi_{xy}(h)$  to represent a covariance between a present value  $x_t$  and  $y_{t-h}$ . Lastly,  $v_x(h)$  represents a covariance between a present value  $x_t$  and  $x_{t-h}$ , for all cases we assume  $h > 0$ .

$$
\gamma_y(0) = \sigma_y^2, \qquad \gamma_y(1) = \theta_1^1 \sigma_y^2, \qquad \gamma_y(2) = \theta_1^2 \sigma_y^2, \n\gamma_y(3) = \theta_1^3 \sigma_y^2, \qquad \gamma_y(4) = \theta_1^4 \sigma_y^2, \qquad \gamma_y(5) = \theta_1^5 \sigma_y^2, \n\gamma_y(6) = \theta_1^6 \sigma_y^2, \qquad \gamma_y(7) = \theta_1^7 \sigma_y^2, \qquad \gamma_y(8) = \theta_1^8 \sigma_y^2, \n\gamma_y(9) = \theta_1^9 \sigma_y^2, \qquad \gamma_y(10) = \theta_1^{10} \sigma_y^2, \n\psi_{yx}(0) = \sigma_{xy}^2, \qquad \psi_{yx}(1) = \theta_1^1 \sigma_{xy}^2, \qquad \psi_{yx}(2) = \theta_1^2 \sigma_{xy}^2, \n\psi_{yx}(3) = \theta_1^3 \sigma_{xy}^2, \qquad \psi_{yx}(4) = \theta_1^4 \sigma_{xy}^2, \qquad \psi_{yx}(5) = \theta_1^5 \sigma_{xy}^2, \n\psi_{yx}(6) = \theta_1^6 \sigma_{xy}^2, \qquad \psi_{yx}(7) = \theta_1^7 \sigma_{xy}^2, \qquad \psi_{yx}(8) = \theta_1^8 \sigma_{xy}^2, \n\psi_{yx}(9) = \theta_1^9 \sigma_{xy}^2, \qquad \psi_{yx}(10) = \theta_1^{10} \sigma_{xy}^2, \n\phi_{xy}(0) = \sigma_{xy}^2, \qquad \psi_{yx}(10) = \theta_1^{10} \sigma_{xy}^2, \n\phi_{xy}(1) = (\lambda_1 \sigma_x^4 + \lambda_2 \sigma_{xy}^2),
$$

$$
\phi_{xy}(1) = (\lambda_1 \sigma_y^4 + \lambda_2 \sigma_{xy}^2), \n\phi_{xy}(2) = (\theta_1^1 \lambda_1 \sigma_y^4 + \lambda_2 \lambda_1 \sigma_y^4 + \lambda_2^2 \sigma_{xy}^2), \n\phi_{xy}(3) = (\theta_1^2 \lambda_1 \sigma_y^4 + \lambda_2 \theta_1^1 \lambda_1 \sigma_y^4 + \lambda_2^2 \lambda_1 \sigma_y^4 + \lambda_2^3 \sigma_{xy}^2), \n\phi_{xy}(4) = (\theta_1^3 \lambda_1 \sigma_y^4 + \lambda_2^4 \sigma_{xy}^2 + \lambda_2^3 \lambda_1 \sigma_y^4 + \lambda_2^2 \theta_1 \lambda_1 \sigma_y^4 + \lambda_2 \theta_1^2 \lambda_1 \sigma_y^4), \n\nu_x(0) = \sigma_x^4, \n\nu_x(1) = (\lambda_1 \sigma_{xy}^2 + \lambda_2 \sigma_x^4), \n\nu_x(2) = (\lambda_1 \theta_1 \sigma_{xy}^2 + \lambda_1 \lambda_2 \sigma_{xy}^2 + \lambda_2^2 \sigma_x^4), \n\nu_x(3) = (\theta_1^2 \lambda_1 \sigma_{xy}^2 + \lambda_2^3 \sigma_x^4 + \lambda_2^2 \lambda_1 \sigma_{xy}^2 + \lambda_2 \lambda_1 \theta_1 \sigma_{xy}^2), \n\nu_x(4) = (\theta_1^3 \lambda_1 \sigma_{xy}^2 + \lambda_2^4 \sigma_x^4 + \lambda_2^3 \lambda_1 \sigma_{xy}^2 + \lambda_2^2 \theta_1 \lambda_1 \sigma_{xy}^2 + \lambda_2 \theta_1^2 \lambda_1 \sigma_{xy}^2)
$$

Consider  $Y_7$ ,  $Y_{7-1}$ , and  $X_7$  to be potentially aggregated time series composed of lagged values of the series  $y_t$  and  $x_t$ , both having a structure represented by Equation [\(3.1\)](#page-11-1). Here,  $Y_{\tau-1}$  represents an aggregation function of lagged values of  $y_t$ . The following results expand upon the estimations presented for simple models in Section [3,](#page-11-0) now utilizing the same aggregation pattern observed in the GDP series.

#### **Example 1**

In this initial example, we assume Case (i),  $\theta_2 = 0$  and  $\lambda_1 = 0$ , i.e.,  $x_t \nrightarrow y_t$  and  $y_t \nightharpoonup x_t$ . Thus,

$$
\begin{array}{rcl}\ny_t &=& \theta_1 y_{t-1} + v_{1,t} \\
x_t &=& \lambda_2 x_{t-1} + v_{2,t} \n\end{array}, \quad v_{j,t} \sim N(0, \sigma^2), \text{ for } j = \{1, 2\} \tag{8.1}
$$

which the aggregation pattern defined by

$$
Y_{\tau} = y_t + 2y_{t-1} + 3y_{t-2} + 2y_{t-3} + y_{t-4},
$$
  
\n
$$
Y_{\tau-1} = y_{t-3} + 2y_{t-4} + 3y_{t-5} + 2y_{t-6} + y_{t-7},
$$
  
\n
$$
X_{\tau} = x_{t-3} + 2x_{t-4} + 3x_{t-5} + 2x_{t-6} + x_{t-7}.
$$
\n(8.2)

Solving for  $\Theta_1$  and  $\Theta_2$ , see equation [\(3.3\)](#page-12-1), we obtain

$$
det(M) = 361\sigma_X^2\sigma_Y^2 + 608\alpha\sigma_X^2\sigma_Y^2 + 380\alpha^2\sigma_X^2\sigma_Y^2 + 152\alpha^3\sigma_X^2\sigma_Y^2 + 38\alpha^4\sigma_X^2\sigma_Y^2 + 608\beta\sigma_X^2\sigma_Y^2 + 380\beta^2\sigma_X^2\sigma_Y^2 + 152\beta^3\sigma_X^2\sigma_Y^2 + 38\beta^4\sigma_X^2\sigma_Y^2 + 1024\alpha\beta\sigma_X^2\sigma_Y^2 + 640\alpha\beta^2\sigma_X^2\sigma_Y^2 + 256\alpha\beta^3\sigma_X^2\sigma_Y^2 + 64\alpha\beta^4\sigma_X^2\sigma_Y^2 + 640\alpha^2\beta\sigma_X^2\sigma_Y^2 + 400\alpha^2\beta^2\sigma_X^2\sigma_Y^2 + 160\alpha^2\beta^3\sigma_X^2\sigma_Y^2 + 40\alpha^2\beta^4\sigma_X^2\sigma_Y^2 + 256\alpha^3\beta\sigma_X^2\sigma_Y^2 + 160\alpha^3\beta^2\sigma_X^2\sigma_Y^2 + 64\alpha^3\beta^3\sigma_X^2\sigma_Y^2 + 16\alpha^3\beta^4\sigma_X^2\sigma_Y^2 + 64\alpha^4\beta\sigma_X^2\sigma_Y^2 + 40\alpha^4\beta^2\sigma_X^2\sigma_Y^2 + 16\alpha^4\beta^3\sigma_X^2\sigma_Y^2 + 4\alpha^4\beta^4\sigma_X^2\sigma_Y^2 + 64\alpha^4\beta\sigma_X^2\sigma_Y^2 + 304\alpha^2\sigma_X^2\sigma_Y^2 + 209\alpha\sigma_X^2\sigma_Y^2 + 76\sigma_X^2\sigma_Y^2 + 304\alpha^4\sigma_X^2\sigma_Y^2 + 190\alpha^5\sigma_X^2\sigma_Y^2 + 76\alpha^6\sigma_X^2\sigma_Y^2 + 209\alpha\sigma_X^2\sigma_Y^2 + 608\alpha^3\beta\sigma_X^2\sigma_Y^2 + 380\alpha^3\beta^2\sigma_X^2\sigma_Y^2 + 152\alpha^3\beta^3\sigma_X^2\sigma_Y^2 + 38\alpha^3\beta^4\sigma_X^2\sigma_Y^2 + 5
$$

#### **Example 2**

Now, we assume Case (iii),  $\theta_2 = 0$  but  $\lambda_1 \neq 0$ , i.e.,  $y_t \to x_t$  but  $x_t \nleftrightarrow y_t$ , or

$$
y_t = \theta_1 y_{t-1} + v_{1,t}
$$
  
\n
$$
x_t = \lambda_1 y_{t-1} + \lambda_2 x_{t-1} + v_{2,t} , \quad v_{j,t} \sim N(0, \sigma^2), \text{ for } j = \{1, 2\}
$$
 (8.3)

with the same aggregation patter defined in the first example, see equation  $(3.5)$ .

$$
det(M) = 361\sigma_x^2\sigma_y^2 + 608\theta_1\sigma_x^2\sigma_y^2 + 380\theta_1^2\sigma_x^2\sigma_y^2 + 152\theta_1^3\sigma_x^2\sigma_y^2 + 38\theta_1^4\sigma_x^2\sigma_y^2 - 361\sigma_x^2y - 608\theta_1\sigma_x^2y - 636\theta_1^2\sigma_x^2y - 472\theta_1^3\sigma_x^2y - 266\theta_1^4\sigma_x^2y - 112\theta_1^5\sigma_x^2y - 36\theta_1^6\sigma_x^2y - 8\theta_1^6\sigma_x^2y - 8\theta_1^6\sigma_x^2y + 608\lambda_2\sigma_x^2\sigma_y^2 + 512\theta_1\lambda_1\sigma_y^2\sigma_{xy} + 380\lambda_x^2\sigma_x^2\sigma_y^2 + 640\theta_1^2\lambda_1\sigma_y^2\sigma_{xy} + 380\lambda_x^2\sigma_x^2\sigma_y^2 + 1024\theta_1\lambda_2\sigma_x^2\sigma_y^2 + 152\theta_1^2\lambda_1\sigma_y^2\sigma_{xy} + 200\theta_1^2\lambda_2\lambda_1\sigma_y^2\sigma_{xy} + 38\lambda_x^4\sigma_x^2\sigma_y^2 + 1024\theta_1\lambda_2\sigma_x^2\sigma_y^2 + 640\theta_1^1\lambda_2^4\sigma_x^2\sigma_y^2 + 256\theta_1\lambda_3^3\sigma_x^2\sigma_y^2 + 24\theta_1^4\lambda_1\sigma_y^2\sigma_{xy} + 64\theta_1^4\lambda_2^4\sigma_x^2\sigma_y^2 + 64\theta_1^2\lambda_2\sigma_x^2\sigma_y^2 + 160\theta_1^3\lambda_2\lambda_1\sigma_y^2\sigma_{xy} + 32\theta_1^3\lambda_2^3\sigma_x^2\sigma_y^2 + 160\theta_1^3\lambda_2^2\sigma_x^2\sigma_y^2 + 160\theta_1^3\lambda_2^2\sigma_x^2\sigma_y^2 + 400\theta_1^3\lambda_2^2\sigma_x^2\sigma_y^2 + 160\theta_1^3\lambda_2^2\sigma_x^2\sigma_y^2 + 4
$$

$$
\begin{array}{ll} \pi_1 &= 361 \theta_1^3 \sigma_x^2 \sigma_y^2 + 304 \theta_1^2 \sigma_x^2 \sigma_y^2 + 209 \theta_1 \sigma_x^2 \sigma_y^2 + 76 \sigma_x^2 \sigma_y^2 + 304 \theta_1^4 \sigma_x^2 \sigma_y^2 \\&+ 190 \theta_1^5 \sigma_x^2 \sigma_y^2 + 76 \theta_1^6 \sigma_x^2 \sigma_y^2 + 19 \theta_1^7 \sigma_x^2 \sigma_y^2 - 733 \theta_1^3 \sigma_x^2 12 \theta_1^4 \sigma_x^2 \\&- 710 \theta_1^5 \sigma_x^2 y - 488 \theta_1^6 \sigma_x^2 y - 266 \theta_1^7 \sigma_x^2 y - 504 \theta_1^2 \sigma_x^2 y - 254 \theta_1 \sigma_x^2 \\&- 76 \sigma_x^2 y - 112 \theta_1^8 \sigma_x^2 y - 36 \theta_1^9 \sigma_x^2 y - 8 \theta_1 0 \sigma_x^2 y - 61 \sigma_x^2 y + 512 \theta_1^3 \lambda_1 \sigma_y^2 \sigma_x \\&+ 608 \theta_1^3 \lambda_2 \sigma_x^2 \sigma_y^2 + 521 \theta_1^4 \lambda_1 \sigma_y^2 \sigma_x y + 265 \theta_1^3 \lambda_2 \lambda_1 \sigma_y^2 \sigma_x y + 380 \theta_1^3 \lambda_2^2 \sigma_x^2 \sigma_y^2 \\&+ 412 \theta_1^5 \lambda_1 \sigma_y^2 \sigma_x y + 252 \theta_1^4 \lambda_2 \lambda_1 \sigma_y^2 \sigma_x y + 280 \theta_1^3 \lambda_2^2 \lambda_1 \sigma_y^2 \sigma_x^2 y \\&+ 247 \theta_1^6 \lambda_1 \sigma_y^2 \sigma_x y + 183 \theta_1^5 \lambda_2 \lambda_1 \sigma_y^2 \sigma_x y + 83 \theta_1^4 \lambda_2^2 \lambda_1 \sigma_y^2 \sigma_x y + 19 \theta_1^3 \lambda_2^3 \lambda_1 \sigma_y^2 \sigma_x \\&+ 38 \theta_1^3 \lambda_2^4 \sigma_x^2 \sigma_y^2 + 382 \theta_1^2 \lambda_1 \sigma_y^2 \sigma_x y + 128 \theta_1^3 \lambda_2^2 \lambda_1 \sigma_y^2 \sigma
$$

$$
\begin{array}{ll} \pi_2 &=& 362\theta_1^3\sigma_y^2\sigma_{xy} + 504\theta_1^4\sigma_y^2\sigma_{xy} + 519\theta_1^5\sigma_y^2\sigma_{xy} + 412\theta_1^6\sigma_y^2\sigma_{xy} + 247\theta_1^7\sigma_y^2\sigma_{xy} \\&+ 184\theta_1^2\sigma_y^2\sigma_{xy} + 45\theta_1\sigma_y^2\sigma_{xy} + 112\theta_1^8\sigma_y^2\sigma_{xy} + 36\theta_1^9\sigma_y^2\sigma_{xy} \\&+ 8\theta_10\sigma_y^2\sigma_{xy} + \theta_11\sigma_y^2\sigma_{xy} - 45\lambda_1\sigma_y^4 - 45\lambda_2\sigma_y^2\sigma_{xy} \\&- 184\theta_1\lambda_1\sigma_y^4 - 144\theta_1\lambda_2\sigma_y^2\sigma_{xy} + - 362\theta_1^2\lambda_1\sigma_y^4 - 236\theta_1^2\lambda_2\sigma_y^2\sigma_{xy} \\&- 504\theta_1^3\lambda_1\sigma_y^4 - 296\theta_1^3\lambda_2\sigma_y^2\sigma_{xy} - 519\theta_1^4\lambda_1\sigma_y^4 - 254\theta_1^4\lambda_2\sigma_y^2\sigma_{xy} \\&- 265\theta_1^3\lambda_2\lambda_1\sigma_y^4 - 190\theta_1^3\lambda_2^2\sigma_y^2\sigma_{xy} - 412\theta_1^5\lambda_1\sigma_y^4 - 183\theta_1^5\lambda_2\lambda_1\sigma_y^4 \\&- 92\theta_1^3\lambda_2^2\lambda_1\sigma_y^4 - 76\theta_1^3\lambda_2^3\sigma_y^2\sigma_{xy} - 247\theta_1^6\lambda_1\sigma_y^4 - 183\theta_1^5\lambda_2\lambda_1\sigma_y^4 \\&- 83\theta_1^4\lambda_2^2\lambda_1\sigma_y^4 - 19\theta_1^3\lambda_2^3\lambda_1\sigma_y^4 - 194\theta_1^3\lambda_2^4\sigma_y^2\sigma_{xy} - 208\theta_1^2\lambda_2\lambda_1\sigma_y^4 \\&- 160\theta_1^2\lambda_
$$

#### **Example 3**

Again, in Case (iii),  $x_t \nrightarrow y_t, y_t \rightarrow x_t$ , model [\(3.6\)](#page-13-1). However, we assume that  $X_\tau$  is non-aggregated, so we are now measuring Granger causality from a non-aggregated to an aggregated series, i.e., which the aggregation pattern defined by

$$
Y_{\tau} = y_t + 2y_{t-1} + 3y_{t-2} + 2y_{t-3} + y_{t-4},
$$
  
\n
$$
Y_{\tau-1} = y_{t-6} + 2y_{t-7} + 3y_{t-8} + 2y_{t-9} + y_{t-10},
$$
  
\n
$$
X_{\tau} = x_{t-2}.
$$
\n(8.4)

The results are given by

$$
det(M) = 19\sigma_x^2 \sigma_y^2 + 32\theta_1 \sigma_x^2 \sigma_y^2 + 20\theta_1^2 \sigma_x^2 \sigma_y^2 + 8\theta_1^3 \sigma_x^2 \sigma_y^2 + 2\theta_1^4 \sigma_x^2 \sigma_y^2 - \theta_1^8 \sigma_{xy}^2 - 4\theta_1^7 \sigma_{xy}^2 - 10\theta_1^6 \sigma_{xy}^2 - 16\theta_1^5 \sigma_{xy}^2 - 19\theta_1^4 \sigma_{xy}^2 - 16\theta_1^3 \sigma_{xy}^2 - 10\theta_1^2 \sigma_{xy}^2 - 4\theta_1 \sigma_{xy}^2 - \sigma_{xy}^2 h \pi_1 = 19\theta_1^3 \sigma_x^2 \sigma_y^2 + 16\theta_1^2 \sigma_x^2 \sigma_y^2 + 11\theta_1 \sigma_x^2 \sigma_y^2 + 4\sigma_x^2 \sigma_y^2 + 16\theta_1^4 \sigma_x^2 \sigma_y^2 + 10\theta_1^5 \sigma_x^2 \sigma_y^2 + 4\theta_1^6 \sigma_x^2 \sigma_y^2 + \theta_1^7 \sigma_x^2 \sigma_y^2 - \theta_1 1 \sigma_{xy}^2 - 4\theta_1 0 \sigma_{xy}^2 - 10\theta_1^9 \sigma_{xy}^2 - 16\theta_1^8 \sigma_{xy}^2 - 19\theta_1^7 \sigma_{xy}^2 - 16\theta_1^6 \sigma_{xy}^2 - 10\theta_1^5 \sigma_{xy}^2 - 4\theta_1^4 \sigma_{xy}^2 - \theta_1^3 \sigma_{xy}^2 \pi_2 = 71\theta_1^7 \sigma_y^2 \sigma_{xy} + 50\theta_1^8 \sigma_y^2 \sigma_{xy} + 21\theta_1^9 \sigma_y^2 \sigma_{xy} + 6\theta_1 0 \sigma_y^2 \sigma_{xy} + \theta_1 1 \sigma_y^2 \sigma_{xy} + 56\theta_1^6 \sigma_y^2 \sigma_{xy} - \theta_1^5 \sigma_y^2 \sigma_{xy} - 58\theta_1^4 \sigma_y^2 \sigma_{xy} - 73\theta_1^3 \sigma_y^2 \sigma_{xy} - 50\theta_1^2 \sigma_y^2 \sigma_{xy} - 19\theta_1 \sigma_y^2 \sigma_{xy} -
$$

As example, testing if further lag values of  $x_t$  are sufficient to reduce considerable spurious effects, we replicate the calculations and simulations mentioned earlier, with  $X_{\tau}$  being non-aggregated and equal to  $x_{t-3}$ , representing one lag further than the previous simulations.

which the aggregation pattern defined by

$$
Y_{\tau} = y_t + 2y_{t-1} + 3y_{t-2} + 2y_{t-3} + y_{t-4},
$$
  
\n
$$
Y_{\tau-1} = y_{t-6} + 2y_{t-7} + 3y_{t-8} + 2y_{t-9} + y_{t-10},
$$
  
\n
$$
X_{\tau} = x_{t-3}.
$$
\n(8.5)

with results given by

$$
det(M) = 19\sigma_x^2 \sigma_y^2 + 32\theta_1 \sigma_x^2 \sigma_y^2 + 20\theta_1^2 \sigma_x^2 \sigma_y^2 + 8\theta_1^3 \sigma_x^2 \sigma_y^2 + 2\theta_1^4 \sigma_x^2 \sigma_y^2 - \theta_1 0 \sigma_{xy}^2 -4\theta_1^9 \sigma_{xy}^2 - 10\theta_1^8 \sigma_{xy}^2 - 16\theta_1^7 \sigma_{xy}^2 - 19\theta_1^6 \sigma_{xy}^2 - 16\theta_1^5 \sigma_{xy}^2 - 10\theta_1^4 \sigma_{xy}^2 -4\theta_1^3 \sigma_{xy}^2 - \theta_1^2 \sigma_{xy}^2 \pi_1 = 19\theta_1^3 \sigma_x^2 \sigma_y^2 + 16\theta_1^2 \sigma_x^2 \sigma_y^2 + 11\theta_1 \sigma_x^2 \sigma_y^2 + 4\sigma_x^2 \sigma_y^2 + 16\theta_1^4 \sigma_x^2 \sigma_y^2 + 10\theta_1^5 \sigma_x^2 \sigma_y^2 +4\theta_1^6 \sigma_x^2 \sigma_y^2 + \theta_1^7 \sigma_x^2 \sigma_y^2 - \theta_1^3 \sigma_{xy}^2 - 4\theta_1^1 2\sigma_{xy}^2 - 10\theta_1 1 \sigma_{xy}^2 - 16\theta_1 0 \sigma_{xy}^2 -19\theta_1^9 \sigma_{xy}^2 - 16\theta_1^8 \sigma_{xy}^2 - 10\theta_1^7 \sigma_{xy}^2 - 4\theta_1^6 \sigma_{xy}^2 - \theta_1^5 \sigma_{xy}^2 \pi_2 = 71\theta_1^8 \sigma_y^2 \sigma_{xy} + 50\theta_1^9 \sigma_y^2 \sigma_{xy} + 21\theta_1 0 \sigma_y^2 \sigma_{xy} + 6\theta_1 1 \sigma_y^2 \sigma_{xy} + \theta_1 2 \sigma_y^2 \sigma_{xy} +56\theta_1^7 \sigma_y^2 \sigma_{xy} - \theta_1^6 \sigma_y^2 \sigma_{xy} - 58\theta_1^5 \sigma_y^2 \sigma_{xy} - 73\theta_1^4 \sigma_y^2 \sigma_{xy} - 50\theta_1^3 \sigma_y^2 \sigma_{xy} -19\theta_1^2 \sigma_y^2 \
$$

#### **Example 4**

Finally, we examine the same causality, Case (iii),  $x_t \nrightarrow y_t, y_t \rightarrow x_t$ , model [\(3.6\)](#page-13-1), but this time, we reverse the aggregation, with  $Y_\tau$  and  $Y_{\tau-1}$  being non-aggregated and  $X_{\tau}$  being aggregated, i.e.,

$$
Y_{\tau} = y_t
$$
  
\n
$$
Y_{\tau-1} = y_{t-1},
$$
  
\n
$$
X_{\tau} = x_{t-6} + 2x_{t-7} + 3x_{t-8} + 2x_{t-9} + x_{t-10}.
$$
\n(8.6)

This allows us to investigate the causality direction from an aggregated series to a non-aggregated one. The results are given by

$$
det(M) = 19\sigma_x^2 \sigma_y^2 - \theta_1^4 \sigma_{xy}^2 - 4\theta_1^5 \sigma_{xy}^2 - 10\theta_1^6 \sigma_{xy}^2 - 16\theta_1^7 \sigma_{xy}^2 - 19\theta_1^8 \sigma_{xy}^2 - 16\theta_1^9 \sigma_{xy}^2
$$
  
\n
$$
-10\theta_1 0 \sigma_{xy}^2 - 4\theta_1 1 \sigma_{xy}^2 - \theta_1 2 \sigma_{xy}^2 + 32\lambda_1 \sigma_y^2 \sigma_{xy} + 32\lambda_2 \sigma_x^2 \sigma_y^2
$$
  
\n
$$
+ 20\theta_1 \lambda_1 \sigma_y^2 \sigma_{xy} + 20\lambda_2 \lambda_1 \sigma_y^2 \sigma_{xy} + 20\lambda_2^2 \sigma_x^2 \sigma_y^2 + 8\theta_1^2 \lambda_1 \sigma_y^2 \sigma_{xy}
$$
  
\n
$$
+ 8\theta_1 \lambda_2 \lambda_1 \sigma_y^2 \sigma_{xy} + 8\lambda_2^2 \lambda_1 \sigma_y^2 \sigma_{xy} + 8\lambda_2^3 \sigma_x^2 \sigma_y^2
$$
  
\n
$$
+ 2\theta_1^3 \lambda_1 \sigma_y^2 \sigma_{xy} + 2\theta_1^2 \lambda_2 \lambda_1 \sigma_y^2 \sigma_{xy} + 2\theta_1 \lambda_2^2 \lambda_1 \sigma_y^2 \sigma_{xy} + 2\lambda_2^3 \lambda_1 \sigma_y^2 \sigma_{xy} + 2\lambda_2^4 \sigma_x^2 \sigma_y^2
$$
  
\n
$$
\pi_1 = 19\theta_1 \sigma_x^2 \sigma_y^2 - \theta_1^5 \sigma_{xy}^2 - 4\theta_1^6 \sigma_{xy}^2 - 10\theta_1^7 \sigma_{xy}^2 - 16\theta_1^8 \sigma_{xy}^2 + 32\theta_1 \lambda_2 \sigma_x^2 \sigma_y^2
$$
  
\n
$$
-10\theta_1 1 \sigma_{xy}^2 - 4\theta_1 2 \sigma_{xy}^2 - \theta_1 3 \sigma_{xy}^2 + 32\theta_1 \lambda_1 \sigma_y^2 \sigma_{xy} + 32\theta_1 \lambda_2 \sigma_x^2 \sigma_y^2
$$
  
\n
$$
+ 20\theta_1^
$$